

# Towards stochastic realization theory for Generalized Linear Switched Systems with inputs: decomposition into stochastic and deterministic components and existence and uniqueness of innovation form

Elie Rouphael, Manas Mejari, Mihaly Petreczky, Lotfi Belkoura

**Abstract**—We study a class of stochastic Generalized Linear Switched System (GLSS), which includes subclasses of jump-Markov, piecewise-linear and Linear Parameter-Varying (LPV) systems. We prove that the output of such systems can be decomposed into deterministic and stochastic components. Using this decomposition, we show existence of state-space representation in innovation form, and we provide sufficient conditions for such representations to be minimal and unique up to isomorphism.

## I. INTRODUCTION

A discrete-time stochastic *Generalized Linear Switched State-Space Representation (GLSS)* is a system

$$\mathbf{S} \begin{cases} \mathbf{x}(t+1) = \sum_{i=1}^{n_\mu} (A_i \mathbf{x}(t) + B_i \mathbf{u}(t) + K_i \mathbf{v}(t)) \boldsymbol{\mu}_i(t) \\ \mathbf{y}(t) = C \mathbf{x}(t) + D \mathbf{u}(t) + F \mathbf{v}(t), \quad t \in \mathbb{Z} \end{cases} \quad (1)$$

where  $A_i \in \mathbb{R}^{n_x \times n_x}$ ,  $B_i \in \mathbb{R}^{n_x \times n_u}$ ,  $K_i \in \mathbb{R}^{n_x \times n_n}$ ,  $i = 1, \dots, n_\mu$ ,  $C \in \mathbb{R}^{n_y \times n_x}$  and  $D \in \mathbb{R}^{n_y \times n_u}$ ,  $F \in \mathbb{R}^{n_y \times n_n}$  are constant matrices. The stochastic processes  $\mathbf{x}$ ,  $\mathbf{u}$ ,  $\mathbf{y}$ ,  $\mathbf{v}$  and  $\boldsymbol{\mu} = (\boldsymbol{\mu}_1, \dots, \boldsymbol{\mu}_{n_\mu})^T$  are the state, input, output, noise and switching processes respectively, taking values respectively in  $\mathbb{R}^{n_x}, \mathbb{R}^{n_u}, \mathbb{R}^{n_y}, \mathbb{R}^{n_n}, \mathbb{R}^{n_\mu}$  and are defined on  $\mathbb{Z}$ .

Intuitively, (1) is a generalization of linear switched systems to the case of infinitely many discrete modes. If  $\boldsymbol{\mu}(t)$  takes values in the set of unit vectors, i.e.,  $\boldsymbol{\mu}_i(t) \in \{0, 1\}$ ,  $i = 1, \dots, n_\mu$ ,  $\sum_{i=1}^{n_\mu} \boldsymbol{\mu}_i(t) = 1$ , then (1) is a *switched system* [15]. If, in addition, the process  $\theta(t)$ , defined by  $\theta(t) = i \iff \boldsymbol{\mu}_i(t) = 1$ , is a Markov chain, then (1) is a jump-Markov system [3]. If  $\boldsymbol{\mu}$  takes values from a possibly infinite set and  $\boldsymbol{\mu}_1 = 1$ , then (1) could be viewed as a subclass of *linear parameter varying (LPV)* systems [17] with an *affine dependence* on scheduling, and  $(\boldsymbol{\mu}_2, \dots, \boldsymbol{\mu}_{n_\mu})^T$  corresponds to the scheduling signal. However, in contrast to LPV systems, in general we are agnostic about the role of  $\boldsymbol{\mu}$  in control, hence the use of the term GLSS.

**Context and motivation:** Consider the following non-stochastic counterpart of (1)

$$\begin{aligned} x(t+1) &= \sum_{i=1}^{n_\mu} (A_i x(t) + B_i u(t) + K_i v(t)) \mu_i(t) \\ y(t) &= C x(t) + D u(t) + F v(t), \quad t \in \mathbb{Z} \end{aligned} \quad (2)$$

where all the signals are deterministic. To emphasize the difference between stochastic and deterministic signals, we use **bold** for the former. Intuitively, (2) is the true system we would like to identify, and (1) is obtained from (2) by considering noise, input and switching signals which are

sampled from the processes  $\mathbf{v}, \mathbf{u}, \boldsymbol{\mu}$  respectively. For the tuple of matrices  $S = (\{A_i, B_i, K_i\}_{i=1}^{n_\mu}, C, D, F)$  of (1) define the *behavior*  $\mathcal{B}_S$  of  $S$  as the set of all tuples of trajectories  $(u, \mu, y)$  such that there exists trajectories  $x$  and  $v$  for which (2) holds. Clearly, all samples paths of  $(\mathbf{u}, \boldsymbol{\mu}, \mathbf{v})$  belong to  $\mathcal{B}_S$ .

The goal of system identification is to find matrices  $\hat{S} = (\{\hat{A}_i, \hat{B}_i, \hat{K}_i\}_{i=1}^{n_\mu}, \hat{C}, \hat{D}, \hat{F})$  from a sample path  $(u, \mu, y)$  of  $(\mathbf{u}, \boldsymbol{\mu}, \mathbf{y})$ , such that the behavior  $\mathcal{B}_{\hat{S}}$  corresponding to the tuple of matrices  $\hat{S}$  is an approximation of  $\mathcal{B}_S$ , i.e., the outputs of the system determined by  $\hat{S}$  should approximate the outputs of (2), for **any** choice of  $u, \mu, v$ , and not only for samples from  $(\mathbf{u}, \boldsymbol{\mu}, \mathbf{v})$ . That is, only the data used for system identification is stochastic, but not the true system itself. This assumption is realistic (measurement error, etc.).

Assuming that the signals used for identification are stochastic allows us to use statistical reasoning about identification algorithms, by viewing  $\hat{S}$  as a statistics for the matrices of (1). In turn, the latter parameterize the joint distribution of  $(\mathbf{u}, \boldsymbol{\mu}, \mathbf{y})$ . Good statistical properties of  $\hat{S}$ , e.g., consistency, guarantee only that the output of the stochastic system determined by  $\hat{S}$  is *close* to that of (1), for the specific stochastic input  $\mathbf{u}$  and switching  $\boldsymbol{\mu}$ . However, this does not imply that the behaviors  $\mathcal{B}_{\hat{S}}$  and  $\mathcal{B}_S$  of  $\hat{S}$  and  $S$  are close. In fact, the outputs of two GLSS may be the same for the input  $\mathbf{u}$  and switching  $\boldsymbol{\mu}$  used during identification, but be different for others [14].

In the LTI case, this gap was resolved by assuming that the data generator (1) and the stochastic system corresponding to  $\hat{S}$  are minimal and in *innovation form*. Since two minimal systems in innovation form with the same outputs and inputs are *isomorphic* [8], if the stochastic system corresponding to  $\hat{S}$  is close to the data generator, then, intuitively, the matrices  $\hat{S}$  and  $S$  are close after a suitable basis transformation, and hence their behaviors are close. Moreover, innovation form is useful for developing and analyzing system identification algorithms, and establishing a correspondence between state-space representations and optimal predictors.

A key step in the proof of existence and uniqueness (up to isomorphism) of minimal LTI systems in innovation form is the decomposition  $\mathbf{y}(t) = \mathbf{y}^d(t) + \mathbf{y}^s(t)$  of the output of the LTI system, such that  $\mathbf{y}^d$  is the output of a noiseless LTI system driven by the input, and  $\mathbf{y}^s$  is the output of an autonomous LTI system which is driven only by the noise. The results on minimal LTI systems in innovation form then follow from realization theory [8] for the autonomous LTI

system generating  $\mathbf{y}^s$ .

**Contribution:** In this paper, we study GLSSs with white noise inputs, uncorrelated with the switching and with the noise. We show that the output  $\mathbf{y}(t)$  of such a GLSS admits the decomposition  $\mathbf{y}(t) = \mathbf{y}^d(t) + \mathbf{y}^s(t)$ , where  $\mathbf{y}^d$  is the output of a GLSS with no noise  $\mathbf{v}$ , and  $\mathbf{y}^s$  is the output of a GLSS with no input  $\mathbf{u}$ . Furthermore, by using results on realization theory of GLSSs with no inputs [11], we use this decomposition to show existence of an innovation form for GLSSs with inputs. Moreover, we present sufficient conditions for minimality and uniqueness (up to isomorphism) of GLSSs in innovation form. Intuitively, an isomorphism between two GLSSs is a linear change of basis, independent of the discrete modes, which transforms the matrices of one GLSS to the corresponding matrices of the other GLSS. This means that minimal GLSSs in innovation form have the same useful properties as their LTI counterparts. In particular, if two minimal GLSSs in innovation form generate (approximately) the same output for the data used for identification, then they generate (approximately) the same output for any input and switching signal.

**Related work:** System identification in general, and subspace methods in particular, of switched [7], jump-Markov [1] and LPV systems [5], [13], [16], [17], [19]–[21] is a well-established topic. Stochastic realization theory of GLSSs with no inputs, i.e., jump-Markov systems with no inputs, bilinear systems and autonomous stochastic LPV systems were addressed in [11]. With respect to [11] the main difference is the presence of the control input  $\mathbf{u}$ .

Existence of a decomposition and existence of innovation form appeared in [9], but only for the case of LPV systems with zero mean i.i.d. scheduling. With respect to [9], the main novelty is that we allow more general switching processes, including finite Markov chains, and that we address minimality and uniqueness of innovation representations. Moreover, in [9] the proofs were not presented.

The existence of innovation representation was studied for LPV systems in [4], [5]. In contrast to this paper, in [4], [5] the noise gain of the innovation representation has a dynamical dependence on the scheduling, and there is no claim on minimality and uniqueness of such representations. In particular, it is unclear when the innovation representation from [4], [5] generates the same output as the original model for all scheduling signals. However, [4], [5] has the advantage that it is valid for any scheduling signal.

## II. PRELIMINARIES

Below we present the notation used in the paper, and we recall a number of concepts from [11] in order to define the subclass of GLSS for which our results hold.

**Probability theory:** We use the terminology of probability theory [2]. The random variables and stochastic processes are understood w.r.t. to a fixed probability space  $(\Omega, \mathcal{F}, P)$ , where  $\mathcal{F}$  is a  $\sigma$ -algebra over the sample space  $\Omega$ , and  $P$  is the probability measure. The expected value of a random variable  $\mathbf{r}$  is denoted by  $E[\mathbf{r}]$  and conditional expectation

w.r.t.  $\sigma$ -algebra  $\mathcal{F}$  is denoted by  $E[\mathbf{r} | \mathcal{F}]$ . The stochastic processes are defined over the discrete time-axis  $\mathbb{Z}$ .

**Finite sequences:** In what follows  $\Sigma$  denotes the finite set (alphabet)  $\Sigma = \{1, \dots, n_\mu\}$ . A *non empty word* over  $\Sigma$  is a finite sequence of letters (elements) of  $\Sigma$ , i.e.,  $w = \sigma_1 \sigma_2 \dots \sigma_k$ , for some  $k \in \mathbb{N}$ ,  $k > 0$ ,  $\sigma_1, \sigma_2, \dots, \sigma_k \in \Sigma$ ;  $|w| := k$  is the length of  $w$ . The set of *all* nonempty words is denoted by  $\Sigma^+$ . We denote the *empty word* by  $\epsilon$  and by convention  $|\epsilon| = 0$ . Let  $\Sigma^* = \{\epsilon\} \cup \Sigma^+$ . The concatenation of two nonempty words  $v = a_1 a_2 \dots a_m$  and  $w = b_1 b_2 \dots b_n$  is defined as  $vw = a_1 \dots a_m b_1 \dots b_n$  for some  $m, n > 0$ . By convention  $v\epsilon = \epsilon v = v$  for all  $v \in \Sigma^*$ .

**Notation for matrices:** We denote by  $I_n$  the  $n \times n$  identity matrix. Consider  $n \times n$  square matrices  $\{A_\sigma\}_{\sigma \in \Sigma}$ . For any  $w = \sigma_1 \sigma_2 \dots \sigma_k \in \Sigma^+$ ,  $k > 0$  and  $\sigma_1, \dots, \sigma_k \in \Sigma$ , we define  $A_w = A_{\sigma_k} A_{\sigma_{k-1}} \dots A_{\sigma_1}$ . For the empty word  $\epsilon$ , set  $A_\epsilon = I_n$ .

**Notions from [11]: admissible switching, Zero Mean Wide Sense Stationary (ZMWSSI), Square Integrable (SII) processes:** We first state our assumptions for the switching process. For every word  $w \in \Sigma^+$  where  $w = \sigma_1 \sigma_2 \dots \sigma_k$ ,  $k \geq 1$ ,  $\sigma_1, \dots, \sigma_k \in \Sigma$ , we define the process  $\mu_w$  as follows:

$$\mu_w(t) = \mu_{\sigma_1}(t - k + 1) \mu_{\sigma_2}(t - k + 2) \dots \mu_{\sigma_k}(t), \quad (3)$$

where  $\mu_i(t)$  is the  $i$ th entry of  $\mu(t)$  for every  $i \in \Sigma$ . For the empty word  $w = \epsilon$ , we set  $\mu_\epsilon(t) = 1$ .

*Definition 1 (Admissible process, [11]):* The switching process  $\mu$  is called *admissible*, if the following holds:

1. There exists a set  $\mathcal{E} \subseteq \Sigma \times \Sigma$  such that:

- $\forall \sigma \in \Sigma, \exists \sigma' \in \Sigma : (\sigma, \sigma') \in \mathcal{E}$ .
- Let  $L$  be the set of all words  $w = \sigma_1 \dots \sigma_k \in \Sigma^+$ ,  $\sigma_1, \dots, \sigma_k \in \Sigma, k > 0$  such that  $(\sigma_i, \sigma_{i+1}) \in \mathcal{E}$ , for all  $i = 1, \dots, k - 1$ . Then for all  $w \in \Sigma^+$ ,  $w \notin L$ ,  $\mu_w = 0$ .

2. Denote by  $\mathcal{F}_t^{\mu, -}$  the  $\sigma$ -algebra generated by the random variables  $\{\mu(k)\}_{k < t}$ . There exists positive numbers  $\{p_\sigma\}_{\sigma \in \Sigma}$  such that for any  $w, v \in \Sigma^+$ ,  $\sigma, \sigma' \in \Sigma$ ,  $t \in \mathbb{Z}$ :

$$E[\mu_{w\sigma}(t) \mu_{v\sigma'}(t) | \mathcal{F}_t^{\mu, -}] = \begin{cases} p_\sigma \mu_w(t-1) \mu_v(t-1) & \sigma = \sigma', \\ & w\sigma \in L \\ & v\sigma \in L \\ 0 & \text{otherwise} \end{cases}$$

$$E[\mu_{w\sigma}(t) \mu_{\sigma'}(t) | \mathcal{F}_t^{\mu, -}] = \begin{cases} p_\sigma \mu_w(t-1) & \sigma = \sigma' \text{ and } \\ & w\sigma \in L \\ 0 & \text{otherwise} \end{cases}$$

$$E[\mu_\sigma(t) \mu_{\sigma'}(t) | \mathcal{F}_t^{\mu, -}] = 0 \quad \text{if } \sigma \neq \sigma'$$

3. There exist real numbers  $\{\alpha_\sigma\}_{\sigma \in \Sigma}$  such that  $\sum_{\sigma \in \Sigma} \alpha_\sigma \mu_\sigma(t) = 1$  for all  $t \in \mathbb{Z}$ .

4. For each  $w, v \in \Sigma^+$ , the process  $[\mu_w, \mu_v]^T$  is square integrable and wide-sense stationary.

From [11] we recall some examples of admissible processes.

*Example 1 (White noise):* If  $\mu = [\mu_1, \mu_2, \dots, \mu_{n_\mu}]^T$  is an i.i.d. process such that  $\mu_1 = 1$ , for all  $i, j = 2, \dots, n_\mu$ ,  $t \in \mathbb{Z}$ ,  $\mu_i(t), \mu_j(t)$  are independent and  $\mu_i(t)$  is zero mean, then  $\mu$  is admissible with  $\mathcal{E} = \Sigma \times \Sigma$  and  $p_\sigma = E[\mu_\sigma^2(t)]$ .

*Example 2 (Discrete valued i.i.d process):* Let  $\theta$  be an i.i.d process with values in  $\Sigma = \{1, \dots, n_\mu\}$ . Let  $\mu_\sigma(t) = \chi(\theta(t) = \sigma)$  for all  $\sigma \in \Sigma$ ,  $t \in \mathbb{Z}$ , where  $\chi$  is the indicator function. Let  $\mathcal{E} = \Sigma \times \Sigma$ , and  $p_\sigma = P(\theta(t) = \sigma)$ ,  $\alpha_\sigma = 1$  for all  $\sigma \in \Sigma$ . Then  $\mu$  is admissible.

*Example 3 (Markov chain):* Let  $\theta$  be a stationary finite state Markov process with values in  $\Theta$ . Assume  $P(\theta(t) = q_2 | \theta(t-1) = q_1) = p_{(q_2, q_1)} > 0$ ,  $q_1, q_2 \in \Theta$ . Let us take  $\Sigma = \Theta \times \Theta$ ,  $\mu_{(q_2, q_1)}(t) = \chi(\theta(t+1) = q_2, \theta(t) = q_1)$  for all  $q_1, q_2 \in \Theta$ ,  $t \in \mathbb{Z}$ , and let  $\mathcal{E} = \{(\sigma_1, \sigma_2) \in \Sigma \times \Sigma | \sigma_1 = (q_2, q_1), \sigma_2 = (q_3, q_2), q_1, q_2, q_3 \in \Theta\}$ . Define  $\alpha_\sigma = 1$  for all  $\sigma \in \Sigma$ . Let us identify  $\Sigma$  with the set  $\{1, \dots, n_\mu\}$ , where  $n_\mu$  is the square of cardinality of  $\Theta$ . Then  $\mu$  is admissible.

*Assumption 1:* The switching process  $\mu$  is admissible. This assumption imposes restrictions on the data used for system identification, but not necessarily for the model class which will be identified. It can be thought of as a persistence of excitation condition. In particular, binary and white noises, which are the simplest persistently exciting signals, satisfy our assumption. Moreover, admissible switching signals cover the fairly general case of Markov chains.

*Remark 1:* For LPV systems, our assumptions imply that the scheduling signal used for identification is stochastic. In addition to this being a persistence of excitation condition, this assumption is also justified by the presence of measurement errors, or when the scheduling is externally generated, or it is a function of the stochastic states/inputs, [13].

Next, we recall the concept of *ZMWSSI process w.r.t  $\mu$*  from [11]. To this end, let  $\{p_\sigma\}_{\sigma \in \Sigma}$  be the constants from Definition 1. For any  $w = \sigma_1 \cdots \sigma_k \in \Sigma^+$ ,  $\sigma_1, \dots, \sigma_k \in \Sigma$ , for a process  $\mathbf{r}$ , define the product  $p_w$  and the process  $\mathbf{z}_w^{\mathbf{r}}$

$$p_w = p_{\sigma_1} p_{\sigma_2} \cdots p_{\sigma_k},$$

$$\mathbf{z}_w^{\mathbf{r}}(t) = \mathbf{r}(t - |w|) \mu_w(t - 1) \frac{1}{\sqrt{p_w}}, \quad (4)$$

where  $\mu_w$  is as in (3). The process  $\mathbf{z}_w^{\mathbf{r}}$  in (4) is interpreted as the product of the *past* of  $\mathbf{r}$  and  $\mu$ .

*Definition 2 (ZMWSSI, [11]):* A process  $\mathbf{r}$  is *Zero Mean Wide Sense Stationary w.r.t.  $\mu$*  (ZMWSSI) if:

(1) For  $t \in \mathbb{Z}$ , the  $\sigma$ -algebras generated by the variables  $\{\mathbf{r}(k)\}_{k \leq t}$ ,  $\{\mu(k)\}_{k < t}$  and  $\{\mu(k)\}_{k \geq t}$ , denoted by  $\mathcal{F}_t^{\mathbf{r}}$ ,  $\mathcal{F}_t^{\mu, -}$  and  $\mathcal{F}_t^{\mu, +}$  respectively, are such that  $\mathcal{F}_t^{\mathbf{r}}$  and  $\mathcal{F}_t^{\mu, +}$  are conditionally independent w.r.t.  $\mathcal{F}_t^{\mu, -}$ .

(2) The processes  $\{\mathbf{r}, \{\mathbf{z}_w^{\mathbf{r}}\}_{w \in \Sigma^+}\}$  are zero mean, square integrable and are jointly wide sense stationary.

Intuitively, ZMWSSI is an extension of wide-sense stationarity, where  $\Sigma^+$  is viewed as time axis: ZMWSSI implies the covariances  $E[\mathbf{z}_w^{\mathbf{r}}(t)(\mathbf{z}_v^{\mathbf{r}}(t))^T]$  do not depend on  $t$ , and they depend on the difference between  $v$  and  $w$ . Next, we recall the definition of a square integrable process w.r.t.  $\mu$ .

*Definition 3 (SII process, [11]):* A process  $\mathbf{r}$  is *Square Integrable w.r.t.  $\mu$*  (SII), if for all  $w \in \Sigma^*$ ,  $t \in \mathbb{Z}$ , the random variable  $\mathbf{r}(t + |w|) \mu_w(t + |w| - 1)$  is square integrable.

As it was mentioned in [11, Section III, Remark 2], if  $\mu$  is essentially bounded, then any ZMWSSI process is SII.

**Assumptions, inputs and outputs and on GLSSs:** First, we define the notion of white noise processes w.r.t.  $\mu$ , which

will be used for stating our assumptions on GLSSs.

*Definition 4 (White noise):* A ZMWSSI process  $\mathbf{r}$  is a white noise w.r.t.  $\mu$ , if for all  $w, v \in \Sigma^*$ ,  $\sigma \in \Sigma$ ,  $\sigma v \in L$ ,

$$E[\mathbf{z}_w^{\mathbf{r}}(t)(\mathbf{z}_{\sigma v}^{\mathbf{r}}(t))^T] = \begin{cases} 0 & \text{if } w \neq \sigma v, w \neq \epsilon \\ E[\mathbf{z}_\sigma^{\mathbf{r}}(t)(\mathbf{z}_\sigma^{\mathbf{r}}(t))^T] & \text{if } w = \sigma v \end{cases}$$

and  $E[\mathbf{r}(t)(\mathbf{z}_{\sigma v}^{\mathbf{r}}(t))^T] = 0$ , and  $E[\mathbf{z}_\sigma^{\mathbf{r}}(t)(\mathbf{z}_\sigma^{\mathbf{r}}(t))^T]$  is nonsingular for all  $\sigma \in \Sigma$

Intuitively, if  $\mathbf{r}$  is a white noise process w.r.t.  $\mu$ , then  $\{\mathbf{z}_w^{\mathbf{r}}(t)\}_{w \in \Sigma^+}$  is a sequence of uncorrelated random variables. Due to 3. of Definition 1, the product  $\mathbf{r}(t - k)\mathbf{r}(t)$  is a linear combination of  $\{\mathbf{z}_w^{\mathbf{r}}(t)\}_{w \in \Sigma^+}$ , hence a white noise process w.r.t.  $\mu$  is also a white noise process in the classical sense. Conversely, if  $\mathbf{r}$  is a white noise and it is independent of  $\{\mu(s)\}_{s \in \mathbb{Z}}$ , then it is a white noise process w.r.t.  $\mu$ .

*Assumption 2 (Inputs and outputs):* (1)  $\mathbf{u}$  is a white noise w.r.t.  $\mu$ , (2) the process  $[\mathbf{y}^T, \mathbf{u}^T]^T$  is a ZMWSSI.

The assumption that  $\mathbf{u}$  is white noise was made for the sake of simplicity, we conjecture that the results of the paper can be extended to more general inputs, e.g., inputs generated by autoregressive models driven by white noise. Next, we define the class of systems considered in this paper.

*Definition 5 (Stationary GLSS):* A *stationary GLSS* (abbreviated sGLSS) of  $(\mathbf{y}, \mathbf{u}, \mu)$  is a system (1), such that

1.  $\mathbf{w} = [\mathbf{v}^T, \mathbf{u}^T]^T$  is a white noise process w.r.t.  $\mu$ , and  $E[\mathbf{v}(t)\mathbf{u}^T(t)\mu_\sigma^2(t)] = 0, \forall \sigma \in \Sigma$ .

2. The process  $[\mathbf{x}^T, \mathbf{w}^T]^T$  is ZMWSSI, and for all  $\sigma \in \Sigma$ ,  $w \in \Sigma^+$ ,  $E[\mathbf{z}_\sigma^{\mathbf{x}}(t)(\mathbf{z}_w^{\mathbf{w}}(t))^T] = 0, E[\mathbf{x}(t)(\mathbf{z}_{\sigma_1}^{\mathbf{w}}(t))^T] = 0$  hold.

3. The eigenvalues of  $\sum_{\sigma \in \Sigma} p_\sigma A_\sigma \otimes A_\sigma$  are inside the open unit circle, where  $\otimes$  denotes the Kronecker product.

4. For all  $\sigma_1, \sigma_2 \in \Sigma$ , if  $(\sigma_1, \sigma_2) \notin \mathcal{E}$ , then  $A_{\sigma_2} A_{\sigma_1} = 0$  and  $A_{\sigma_2} [B_{\sigma_1}, K_{\sigma_1}] E[\mathbf{z}_{\sigma_1}^{\mathbf{w}}(t)(\mathbf{z}_{\sigma_1}^{\mathbf{w}}(t))^T] = 0$ .

If  $B_i = 0, i \in \Sigma$ , and  $D = 0$  then we call (1) an *autonomous stationary GLSS* (asGLSS) of  $(\mathbf{y}, \mu)$ .

*Remark 2:* In stochastic realization theory of LTI systems [12], it is assumed that  $\mathbf{y}$  does not Granger cause  $\mathbf{u}$ . Note that part (1) of Definition 2, when applied to  $\mathbf{y}$ , is similar to Granger causality, but it captures absence of feedback from  $\mathbf{y}$  to  $\mu$ , hence it is different from Granger-causality in [12].

*Remark 3:* Note that  $\mathbf{u}$  and  $\mathbf{v}$  may have different variances, as long as they satisfy condition 1. of Definition 5.

Intuitively, sGLSSs are introduced in order to ensure that all the relevant stochastic processes are stationary in a suitable sense. The latter is a widespread assumption in system identification. In the terminology of [11], a sGLSS (resp. asGLSS) corresponds to a stationary *Generalized Bilinear System (GBS)* with noise  $[\mathbf{v}^T, \mathbf{u}^T]^T$  (resp.  $\mathbf{v}$ ).

From [11] it follows that the processes  $\mathbf{x}$  and  $\mathbf{y}$  are ZMWSSI, and hence Assumption 2 is satisfied for all  $(\mathbf{u}, \mu, \mathbf{y})$  generated by sGLSS. Indeed, by [11, Lemma 2]

$$\mathbf{x}(t) = \sum_{\substack{\sigma \in \Sigma, w \in \Sigma^* \\ \sigma w \in L}} \sqrt{p_{\sigma w}} A_w \left( K_\sigma \mathbf{z}_{\sigma w}^{\mathbf{v}}(t) + B_\sigma \mathbf{z}_{\sigma w}^{\mathbf{u}}(t) \right), \quad (5)$$

where the infinite sum converges in the mean square sense. Hence, the state  $\mathbf{x}$  is uniquely determined by the system matrices and the input  $\mathbf{u}$  and noise  $\mathbf{v}$ , and it is the limit of any state trajectory started from some initial state.

*Notation 1:* We identify the sGLSS  $\mathbf{S}$  of the form (1) with the tuple  $\mathbf{S} = (\{A_\sigma, K_\sigma, B_\sigma\}_{\sigma=1}^{n_\mu}, C, D, F, \mathbf{v})$ , and if  $\mathbf{S}$  is a asGLSS, i.e.  $B_\sigma = 0$ ,  $\sigma \in \Sigma$ ,  $D = 0$ , then we will identify it with the tuple  $\mathbf{S} = (\{A_\sigma, K_\sigma\}_{\sigma=1}^{n_\mu}, C, F, \mathbf{v})$ .

### III. MAIN RESULT

We start by recalling from [11] the following terminology.

*Notation 2 (Orthogonal projection  $E_l$ ):* Recall from [2] that the set of real valued square integrable random variables, denoted by  $\mathcal{H}_1$ , forms a Hilbert-space with the scalar product defined as  $\langle \mathbf{z}_1, \mathbf{z}_2 \rangle = E[\mathbf{z}_1 \mathbf{z}_2]$ . Let  $\mathbf{z}$  be a square integrable random variable taking its values in  $\mathbb{R}^k$ . Let  $M$  be a closed subspace of  $\mathcal{H}_1$ . The orthogonal projection of  $\mathbf{z}$  onto  $M$ , denoted by  $E_l[\mathbf{z} \mid M]$ , is defined as the random variable  $\mathbf{z}^* = [\mathbf{z}_1^*, \dots, \mathbf{z}_k^*]^T$  such that  $\mathbf{z}_i^* \in M$  is the orthogonal projection of the  $i$ th coordinate  $\mathbf{z}_i$  of  $\mathbf{z}$ , viewed as an element of  $\mathcal{H}_1$  onto  $M$ . If  $\mathfrak{S}$  is a subset of square integrable random variables in  $\mathbb{R}^p$ , and  $M$  is generated by the coordinates of the elements of  $\mathfrak{S}$ , then instead of  $E_l[\mathbf{z} \mid M]$  we use  $E_l[\mathbf{z} \mid \mathfrak{S}]$ . Intuitively,  $E_l[\mathbf{z} \mid \mathfrak{S}]$  is the best (minimal variance) linear prediction of  $\mathbf{z}$  based on the elements of  $\mathfrak{S}$ .

Using the notation above, let us define the *deterministic component*  $\mathbf{y}^d$  of  $\mathbf{y}$  as follows

$$\mathbf{y}^d(t) = E_l[\mathbf{y}(t) \mid \{\mathbf{z}_w^u(t)\}_{w \in \Sigma^+} \cup \{\mathbf{u}(t)\}]. \quad (6)$$

Also, define the *stochastic component* of  $\mathbf{y}$  as

$$\mathbf{y}^s(t) = \mathbf{y}(t) - \mathbf{y}^d(t). \quad (7)$$

Intuitively,  $\mathbf{y}^d(t)$  represents the best prediction of  $\mathbf{y}(t)$  which is linear in the present and past values of  $\mathbf{u}$  and non-linear in the past values of  $\boldsymbol{\mu}$ . In fact,  $\mathbf{y}^d$  is the output of the asGLSS obtained from (1) by considering  $\mathbf{v} = 0$  and viewing  $\mathbf{u}$  as noise. That is,  $\mathbf{y}^d(t)$  is the output of a system with no noise, hence  $\mathbf{y}^d(t)$  is a deterministic function of  $\{\mathbf{u}(s), \boldsymbol{\mu}(s)\}_{s < t} \cup \{\mathbf{u}(t)\}$ . This motivates us to call  $\mathbf{y}^d$  the deterministic component, similarly to LTI literature [6, Definition 9.3]. In contrast,  $\mathbf{y}^s(t)$  is the output of the asGLSS obtained from (1) by taking  $\mathbf{u} = 0$  and viewing  $\mathbf{v}$  as noise, and thus  $\mathbf{y}^s$  does not depend on  $\mathbf{u}$ .

*Theorem 1:* For a sGLSS of the form (1),  $\mathbf{S}_d = (\{A_\sigma, B_\sigma\}_{\sigma=1}^{n_\mu}, C, D, \mathbf{u})$  is an asGLSS of  $(\mathbf{y}^d, \boldsymbol{\mu})$  and  $\mathbf{S}_s = (\{A_\sigma, K_\sigma\}_{\sigma=1}^{n_\mu}, C, F, \mathbf{v})$  is an asGLSS of  $(\mathbf{y}^s, \boldsymbol{\mu})$ .

The proof of Theorem 1 is presented in the Appendix A.

The converse of Theorem 1 also holds. To this end, recall from [11] the definition of the *innovation process* of  $\mathbf{y}^s$ :

$$\mathbf{e}^s(t) = \mathbf{y}^s(t) - E_l[\mathbf{y}^s(t) \mid \{\mathbf{z}_w^{\mathbf{y}^s}(t)\}_{w \in \Sigma^+}] \quad (8)$$

Intuitively,  $\mathbf{e}^s(t)$  is the difference between  $\mathbf{y}^s(t)$  and the best linear prediction of  $\mathbf{y}^s(t)$  based on its own past multiplied with past values of the switching process.

*Theorem 2:* Assume that there exists a sGLSS of  $(\mathbf{y}, \mathbf{u}, \boldsymbol{\mu})$  and that the following holds:

1.  $\hat{\mathbf{S}}_d = (\{\hat{A}_i, \hat{B}_i\}_{i=1}^{n_\mu}, \hat{C}, \hat{D}, \mathbf{u})$  is an asGLSS of  $(\mathbf{y}^d, \boldsymbol{\mu})$ .
2.  $\hat{\mathbf{S}}_s = (\{\hat{A}_i, \hat{K}_i\}_{i=1}^{n_\mu}, \hat{C}, I_{n_y}, \mathbf{v})$  is an asGLSS of  $(\mathbf{y}^s, \boldsymbol{\mu})$  in innovation form, i.e.  $\mathbf{v} = \mathbf{e}^s$ .

Then  $\hat{\mathbf{S}} = (\{\hat{A}_i, \hat{K}_i, \hat{B}_i\}_{i=1}^{n_\mu}, \hat{C}, \hat{D}, I, \mathbf{e}^s)$  is a sGLSS of  $(\mathbf{y}, \mathbf{u}, \boldsymbol{\mu})$ , and  $\mathbf{e}^s(t) = \mathbf{y}(t) - \hat{\mathbf{y}}(t)$ , where

$$\hat{\mathbf{y}}(t) = E_l[\mathbf{y}(t) \mid \{\mathbf{z}_w^{\mathbf{y}}(t), \mathbf{z}_w^{\mathbf{u}}(t)\}_{w \in \Sigma^+} \cup \{\mathbf{u}(t)\}]. \quad (9)$$

The proof of Theorem 2 is presented in Section B.

*Remark 4:* The condition that the matrices  $\{A_i\}_{i=1}^{n_\mu}$  and  $C$  of  $\hat{\mathbf{S}}_d$  and of  $\hat{\mathbf{S}}_s$  are the same can be relaxed: if  $\hat{\mathbf{S}}_d = (\{\hat{A}_i^d, \hat{B}_i^d\}_{i=1}^{n_\mu}, \hat{C}^d, \hat{D}, \mathbf{u})$  and  $\hat{\mathbf{S}}_s = (\{\hat{A}_i^s, \hat{B}_i^s\}_{i=1}^{n_\mu}, \hat{C}^s, I, \mathbf{e}^s)$  are asGLSS of  $(\mathbf{y}^s, \boldsymbol{\mu})$  and  $(\mathbf{y}^s, \boldsymbol{\mu})$  respectively, then with

$$\hat{A}_i = \begin{bmatrix} \hat{A}_i^d & 0 \\ 0 & \hat{A}_i^s \end{bmatrix}, \quad \hat{B}_i = \begin{bmatrix} \hat{B}_i^d \\ 0 \end{bmatrix}, \quad \hat{K}_i = \begin{bmatrix} 0 \\ \hat{K}_i^s \end{bmatrix}, \quad \hat{C} = \begin{bmatrix} (C^d)^T \\ (C^s)^T \end{bmatrix}^T,$$

$\hat{\mathbf{S}}_d$  and  $\hat{\mathbf{S}}_s$  from Theorem 2 are asGLSSs of  $(\mathbf{y}^d, \boldsymbol{\mu})$  and  $(\mathbf{y}^s, \boldsymbol{\mu})$  respectively and Theorem 2 applies.

Thus, Theorem 1 – 2 means that finding sGLSSs of  $(\mathbf{y}, \mathbf{u}, \boldsymbol{\mu})$  is equivalent to finding an asGLSS representations of the deterministic and stochastic components respectively.

Theorem 2 suggests that  $\mathbf{e}^s(t)$  can be viewed as the innovation process of  $\mathbf{y}$ . Indeed,  $\hat{\mathbf{y}}(t)$  from (9) is the best linear prediction of  $\mathbf{y}(t)$  based on past values of  $\mathbf{y}$  and past and current values of  $\mathbf{u}$  multiplied by past values of the switching process. Then  $\mathbf{e}^s(t)$  is the prediction error  $\mathbf{y}(t) - \hat{\mathbf{y}}(t)$ . This motivates the following definition.

*Definition 6 (Innovation form):* The sGLSS (1) is in *innovation form*, if  $F$  is the identity matrix and  $\mathbf{v} = \mathbf{e}^s$ .

Similarly to the LTI case [8], an sGLSS in innovation form can be viewed as a recursive filter driven by  $\mathbf{y}$ ,  $\mathbf{u}$ ,  $\boldsymbol{\mu}$ , whose output is the optimal prediction  $\hat{\mathbf{y}}(t)$  from (9). Indeed, from  $\mathbf{e}^s(t) = \mathbf{y}(t) - C\mathbf{x}(t) - D\mathbf{u}(t)$  it follows that  $\mathbf{x}(t+1)$  is a function of  $\mathbf{x}(t)$ ,  $\mathbf{u}(t)$ ,  $\mathbf{y}(t)$ ,  $\boldsymbol{\mu}(t)$ , and  $\hat{\mathbf{y}}(t) = C\mathbf{x}(t) + D\mathbf{u}(t)$ .

In what follows we make the assumption that  $\mathbf{y}^s$  is SII, which is always satisfied when  $\boldsymbol{\mu}$  is bounded<sup>1</sup>.

*Corollary 1 (Existence):* Assume that  $\mathbf{y}^s$  is SII. Then any sGLSS of  $(\mathbf{y}, \mathbf{u}, \boldsymbol{\mu})$  can be transformed to a sGLSS of  $(\mathbf{y}, \mathbf{u}, \boldsymbol{\mu})$  in innovation form.

*Proof:* From Theorem 1 it follows that  $\mathbf{S}_s$  is an asGLSS of  $(\mathbf{y}^s, \boldsymbol{\mu})$  and  $\mathbf{S}_d$  is an asGLSS of  $(\mathbf{y}^d, \boldsymbol{\mu})$ . From [11, Theorem 2] it follows that there exists a (minimal) asGLSS  $\hat{\mathbf{S}}_s$  of  $(\mathbf{y}^s, \boldsymbol{\mu})$  in innovation form and by [11, Theorem 3] it can be computed from  $\mathbf{S}_s$  using [11, Algorithm 1]. Then using Remark 4 and Theorem 2, it follows that  $\hat{\mathbf{S}}$  defined in Theorem 2 is a sGLSS of  $(\mathbf{y}, \mathbf{u}, \boldsymbol{\mu})$  in innovation form. ■ Next, we formulate conditions for minimality of sGLSSs. To this end, we define the *dimension* the sGLSS  $\mathbf{S}$  from (1), denoted by  $\dim(\mathbf{S})$ , as the dimension  $n_x$  of its state-space.

*Corollary 2 (Minimality):* Assume that  $\mathbf{S}$  is a sGLSS of  $(\mathbf{y}, \mathbf{u}, \boldsymbol{\mu})$ ,  $\mathbf{y}^s$  is SII, and assume that  $\mathbf{S}_s$  from Theorem 1 is observable and reachable in the terminology of [11], if viewed as a stationary GBS. Then the dimension of  $\mathbf{S}$  is minimal among all the sGLSSs of  $(\mathbf{y}, \mathbf{u}, \boldsymbol{\mu})$ .

*Proof:* Let  $\hat{\mathbf{S}}$  be a sGLSS of  $(\mathbf{y}, \mathbf{u}, \boldsymbol{\mu})$  such that  $\dim(\hat{\mathbf{S}}) < \dim(\mathbf{S})$ . Then by Theorem 1,  $\hat{\mathbf{S}}_s$  is an asGLSS of  $(\mathbf{y}^s, \boldsymbol{\mu})$  of the same dimension as  $\hat{\mathbf{S}}$ . However, from [11, Theorem 2],  $\mathbf{S}_s$  is a minimal dimensional asGLSS of  $(\mathbf{y}^s, \boldsymbol{\mu})$  and hence  $\dim(\mathbf{S}_s) \leq \dim(\hat{\mathbf{S}}_s)$ . However,  $\dim(\mathbf{S}_s) = \dim(\mathbf{S})$  by construction and hence  $\dim(\mathbf{S}) \leq \dim(\hat{\mathbf{S}})$ , which is a contradiction. ■

Note that observability and reachability in the sense of [11] can be characterized by rank conditions of suitable matrices,

<sup>1</sup>This follows from [11, Remark 2] and from the fact that  $\mathbf{y}^s$  is the output of a asGLSS, and thus by [11] it is ZMWSSI.

which can be constructed from the matrices of  $\mathbf{S}_s$ . We also get the following sufficient condition for isomorphism of minimal sGLSSs in innovation form.

*Corollary 3 (Isomorphism):* Assume that  $\mathbf{S}$  is of the form (1) and  $\hat{\mathbf{S}} = (\{\hat{A}_\sigma, \hat{B}_\sigma, \hat{K}_\sigma\}_{\sigma \in \Sigma}, \hat{C}, \hat{D}, I, \mathbf{e}^s)$  and they are both sGLSS of  $(\mathbf{y}, \mathbf{u}, \boldsymbol{\mu})$  in innovation form and  $\mathbf{S}_s$  and  $\hat{\mathbf{S}}_s$  are both reachable and observable as stationary GBS in the terminology of [11]. Assume that  $\mathbf{y}^s$  is SII, and that the covariance matrix  $E[\mathbf{e}^s(t)(\mathbf{e}^s(t))^T \boldsymbol{\mu}_\sigma^2(t)]$  is nonsingular and  $\text{Im}[B_\sigma^T, \hat{B}_\sigma^T]^T \subseteq \text{Im}[K_\sigma^T, \hat{K}_\sigma^T]^T$ ,  $\sigma \in \Sigma$  and  $\hat{D} = D$ . Then there exists a non-singular matrix  $T$  such that for all  $\sigma \in \Sigma$ ,

$$TA_\sigma T^{-1} = \hat{A}_\sigma, T[K_\sigma, B_\sigma] = [\hat{K}_\sigma, \hat{B}_\sigma], CT^{-1} = \hat{C} \quad (10)$$

*Proof:* Since  $\mathbf{S}_s$  and  $\hat{\mathbf{S}}_s$  are both minimal asGLSS of  $(\mathbf{y}, \boldsymbol{\mu})$  in innovation form, and by [11, Theorem 2], they are isomorphic, i.e., there exists a non-singular matrix  $T$  such that  $TA_\sigma T^{-1} = \hat{A}_\sigma$ ,  $TK_\sigma = \hat{K}_\sigma$ ,  $CT^{-1} = \hat{C}$  holds. Since  $\text{Im}[B_\sigma^T, \hat{B}_\sigma^T]^T \subseteq \text{Im}[K_\sigma^T, \hat{K}_\sigma^T]^T$ , for some matrix  $Z_\sigma$ ,  $B_\sigma = K_\sigma Z_\sigma$ ,  $\hat{B}_\sigma = \hat{K}_\sigma Z_\sigma$ , from which (10) follows. ■

#### IV. CONCLUSION

We have shown that outputs of stochastic generalized linear switched systems can be decomposed into two parts, deterministic and stochastic one, and we used it to derive existence of representation in innovation form and to formulate sufficient conditions for minimality and uniqueness of such representations up to isomorphism. Future work will be directed towards extending these results for a larger class of inputs and switching signals.

#### APPENDIX

##### PROOFS OF THEOREMS 1 AND 2

In what follows, for a ZMWSSI process  $\mathbf{r}$ , we denote by  $\mathcal{H}_{t,+}^{\mathbf{r}}$  the closed subspace of  $\mathcal{H}_1$  (see Notation 2) generated by the entries of  $\{\mathbf{z}_w^{\mathbf{r}}(t)\}_{w \in \Sigma^+} \cup \{\mathbf{r}(t)\}$  and by  $\mathcal{H}_t^{\mathbf{r}}$  the closed subspace of  $\mathcal{H}_1$  generated by the entries of  $\{\mathbf{z}_w^{\mathbf{r}}(t)\}_{w \in \Sigma^+}$ .

##### A. Proof of Theorem 1

*Lemma 1:* The entries of  $\mathbf{v}(t)$ ,  $\{\mathbf{z}_w^{\mathbf{v}}(t)\}_{w \in \Sigma^+}$  are orthogonal to  $\mathcal{H}_{t,+}^{\mathbf{u}}$ .

*Proof:* By definition,  $\mathbf{w}(t) = [\mathbf{v}^T(t), \mathbf{u}^T(t)]^T$  is a white noise process w.r.t.  $\boldsymbol{\mu}$ , and  $\mathbf{v}(t)$  is the upper  $n_n$  block of  $\mathbf{w}(t)$ . Since  $\mathbf{w}$  is a white noise w.r.t.  $\boldsymbol{\mu}$ ,  $E[\mathbf{w}(t)(\mathbf{z}_w^{\mathbf{w}}(t))^T] = 0$ , and  $\frac{1}{\sqrt{p_\sigma}} E[\mathbf{v}(t)(\mathbf{z}_w^{\mathbf{u}}(t))^T]$  is the upper-left block of that latter matrix, and hence it is also zero. From the definition of sGLSS, it follows that  $E[\mathbf{v}(t)(\mathbf{u}(t))^T \boldsymbol{\mu}_i^2(t)] = 0$ ,  $i \in \Sigma$ . Since  $E[\mathbf{v}(t)(\mathbf{u}(t))^T \boldsymbol{\mu}_i(t) \boldsymbol{\mu}_j(t)] = 0$  for  $i \neq j$  due to  $\mathbf{w}$  being ZMWSSI ([11, Lemma 7]), and  $\sum_{i=1}^{n_p} \alpha_i \boldsymbol{\mu}_i = 1$  for some  $\{\alpha_i\}_{i=1}^{n_p}$ , it follows  $E[\mathbf{v}(t)(\mathbf{u}(t))^T] = \sum_{i,j=1}^{n_p} \alpha_i \alpha_j E[\mathbf{v}(t)(\mathbf{u}(t))^T \boldsymbol{\mu}_i(t) \boldsymbol{\mu}_j(t)] = 0$ . That is,  $\mathbf{v}(t)$  is orthogonal to  $\mathcal{H}_{t,+}^{\mathbf{u}}$ . Since  $\mathbf{w}(t)$  is a ZMWSSI, from [11, Lemma 7] it follows that  $E[\mathbf{z}_w^{\mathbf{w}}(t)(\mathbf{z}_v^{\mathbf{w}}(t))^T] = 0$  for all  $v \in \Sigma^+$ ,  $v \neq w$  or  $v \notin L$  or  $w \notin L$ , and if  $v = w \in L$  and  $\sigma$  is the first letter of  $w$ , then  $E[\mathbf{z}_w^{\mathbf{w}}(t)(\mathbf{z}_w^{\mathbf{w}}(t))^T] = E[\mathbf{z}_\sigma^{\mathbf{w}}(t)(\mathbf{z}_\sigma^{\mathbf{w}}(t))^T]$ . Since  $E[\mathbf{z}_w^{\mathbf{v}}(t)(\mathbf{z}_v^{\mathbf{u}}(t))^T]$  is the upper right block of  $E[\mathbf{z}_w^{\mathbf{w}}(t)(\mathbf{z}_v^{\mathbf{w}}(t))^T]$ , it follows that  $E[\mathbf{z}_w^{\mathbf{v}}(t)(\mathbf{z}_v^{\mathbf{u}}(t))^T] = 0$  if  $v \neq w$ , or either  $v \notin L$  or  $w \notin L$ , and  $E[\mathbf{z}_w^{\mathbf{v}}(t)(\mathbf{z}_w^{\mathbf{u}}(t))^T] =$

$E[\mathbf{z}_\sigma^{\mathbf{v}}(t)(\mathbf{z}_\sigma^{\mathbf{u}}(t))^T] = \frac{1}{p_\sigma} E[\mathbf{u}(t-1)\mathbf{v}(t-1)\boldsymbol{\mu}_\sigma^2(t-1)]$  for any  $w \in L$ , where  $\sigma$  is the first letter of  $w$ . From Definition 5, it follows that the latter expectation is zero. That is,  $E[\mathbf{z}_w^{\mathbf{v}}(t)(\mathbf{z}_v^{\mathbf{u}}(t))^T] = 0$  for all  $v \in \Sigma^+$ .

Since  $\mathbf{w}(t)$  is a white noise w.r.t.  $\boldsymbol{\mu}$ , by [11, Lemma 7]  $E[\mathbf{z}_w^{\mathbf{w}}(t)(\mathbf{w}(t))^T] = E[\mathbf{z}_{ws}^{\mathbf{w}}(t)(\mathbf{z}_s^{\mathbf{w}}(t))^T] = 0$  for any  $s \in \Sigma^+$ , and since  $E[\mathbf{z}_w^{\mathbf{v}}(t)(\mathbf{u}(t))^T]$  is the upper right block of  $E[\mathbf{z}_w^{\mathbf{w}}(t)(\mathbf{w}(t))^T]$ ,  $E[\mathbf{z}_w^{\mathbf{v}}(t)(\mathbf{u}(t))^T] = 0$ . Since  $\mathbf{z}_w^{\mathbf{v}}(t)$  is orthogonal to the generators of  $\mathcal{H}_{t,+}^{\mathbf{u}}$ , the lemma follows. ■

*Lemma 2:* Define  $\mathbf{x}^d(t) = E[\mathbf{x}(t) | \mathcal{H}_{t,+}^{\mathbf{u}}]$ . The entries of  $\mathbf{x}^d(t)$  belong to  $\mathcal{H}_t^{\mathbf{u}}$  and

$$\mathbf{x}^d(t) = \sum_{w \in \Sigma^*, \sigma \in \Sigma, \sigma w \in L} \sqrt{p_{\sigma w}} A_w B_\sigma \mathbf{z}_{\sigma w}^{\mathbf{u}}(t), \quad (11)$$

where the convergence is in the mean square sense.

*Proof:* It is clear from the definition that the components of  $\mathbf{x}^d(t)$  belong to  $\mathcal{H}_{t,+}^{\mathbf{u}}$ . From Lemma 1 it follows that,  $E[\mathbf{z}_{\sigma w}^{\mathbf{v}}(t) | \mathcal{H}_{t,+}^{\mathbf{u}}] = 0$ , and since the components of  $\mathbf{z}_{\sigma w}^{\mathbf{u}}(t)$  belong to  $\mathcal{H}_{t,+}^{\mathbf{u}}$ , it follows that  $E[\mathbf{z}_{\sigma w}^{\mathbf{u}}(t) | \mathcal{H}_{t,+}^{\mathbf{u}}] = \mathbf{z}_{\sigma w}^{\mathbf{u}}(t)$ . Since (5) holds and the map  $z \mapsto E[z | M]$  (where  $z \in \mathcal{H}_1$ ) is a continuous linear operator for any closed subspace  $M$ , it follows that  $\mathbf{x}^d(t)$  will be the infinite sum of the elements  $\sqrt{p_{\sigma w}} A_w (K_\sigma E[\mathbf{z}_{\sigma w}^{\mathbf{v}}(t) | \mathcal{H}_{t,+}^{\mathbf{u}}] + B_\sigma E[\mathbf{z}_{\sigma w}^{\mathbf{u}}(t) | \mathcal{H}_{t,+}^{\mathbf{u}}])$ , i.e., (11) holds. Since the components of  $\mathbf{z}_{\sigma w}^{\mathbf{u}}(t)$  belong to  $\mathcal{H}_t^{\mathbf{u}}$ , the components of the right-hand side of (11) belong to  $\mathcal{H}_t^{\mathbf{u}}$  and hence the components of  $\mathbf{x}^d(t)$  belong to  $\mathcal{H}_t^{\mathbf{u}}$ . The convergence of the right-hand side of (11) in the mean square sense follows from that of (5). ■

*Lemma 3:* Define  $\mathbf{x}^s(t) = \mathbf{x}(t) - \mathbf{x}^d(t)$ . The entries of  $\mathbf{x}^s(t)$  belong to  $\mathcal{H}_t^{\mathbf{y}}$ , they are orthogonal to  $\mathcal{H}_{t,+}^{\mathbf{u}}$  and

$$\mathbf{x}^s(t) = \sum_{w \in \Sigma^*, \sigma \in \Sigma, \sigma w \in L} \sqrt{p_{\sigma w}} A_w K_\sigma \mathbf{z}_{\sigma w}^{\mathbf{v}}(t), \quad (12)$$

where the sum converges in the mean-square sense.

*Proof:* From (11),  $\mathbf{x}^s(t) = \mathbf{x}(t) - \mathbf{x}^d(t)$  and (5), it follows that (12) holds. By Lemma 1,  $\{\mathbf{z}_w^{\mathbf{v}}(t)\}_{w \in \Sigma^+}$  are orthogonal to  $\mathcal{H}_{t,+}^{\mathbf{u}}$ , hence all the summands in the right-hand side of (12) are orthogonal to  $\mathcal{H}_{t,+}^{\mathbf{u}}$ . ■

*Proof:* (Proof of Theorem 1): The proof is an extension of [10, proof of Lemma 1]. Since the eigenvalues of  $\sum_{\sigma \in \Sigma} p_\sigma A_\sigma \otimes A_\sigma$  are inside the open unit disk and  $\mathbf{u}$  and  $\mathbf{v}$  are white noise processes, from (11)-(12) and [11, Lemma 3] it follows that  $\mathbf{x}^d$  is the unique state process of  $\mathbf{S}_d$  and  $\mathbf{x}^s$  is the unique state process of  $\mathbf{S}_s$ . Notice that  $\mathbf{y}^d(t) = C E[\mathbf{x}(t) | \mathcal{H}_{t,+}^{\mathbf{u}}] + D E[\mathbf{u}(t) | \mathcal{H}_{t,+}^{\mathbf{u}}] + E[\mathbf{v}(t) | \mathcal{H}_{t,+}^{\mathbf{u}}]$ . By Lemma 1,  $\mathbf{v}(t)$  is orthogonal to  $\mathcal{H}_{t,+}^{\mathbf{u}}$ ,  $E[\mathbf{v}(t) | \mathcal{H}_{t,+}^{\mathbf{u}}] = 0$  and as the components  $\mathbf{u}(t)$  belong to  $\mathcal{H}_{t,+}^{\mathbf{u}}$ ,  $E[\mathbf{u}(t) | \mathcal{H}_{t,+}^{\mathbf{u}}] = \mathbf{u}(t)$ . Hence,  $\mathbf{y}^d(t) = C \mathbf{x}^d(t) + D \mathbf{u}(t)$  and  $\mathbf{y}^s(t) = C \mathbf{x}^s(t) + F \mathbf{v}(t)$ . That is,  $\mathbf{S}_d$  is an asGLSS of  $(\mathbf{y}^d, \boldsymbol{\mu})$  and  $\mathbf{S}_s$  is an asGLSS of  $(\mathbf{y}^s, \boldsymbol{\mu})$  respectively. ■

##### B. Proof of Theorem 2

Assume that  $\mathbf{S}$  of the form (1) is a sGLSS of  $(\mathbf{y}, \mathbf{u}, \boldsymbol{\mu})$ .

*Lemma 4:* The entries of the random variables  $\mathbf{y}^s(t)$ ,  $\mathbf{e}^s(t)$ ,  $\{\mathbf{z}_v^{\mathbf{y}^s}(t), \mathbf{z}_v^{\mathbf{e}^s}(t)\}_{v \in \Sigma^+}$  belong to  $\mathcal{H}_{t,+}^{\mathbf{y}}$ .

*Proof:* Recall from the proof of Theorem 1 that  $\mathbf{y}^s(t) = C \mathbf{x}^s(t) + \mathbf{v}(t)$ . Then by (12), the components of  $\mathbf{y}^s(t)$

belong to  $\mathcal{H}_{t,+}^{\mathbf{y}}$ . Then by [11, Lemma 11], the components of  $\mathbf{z}_v^{\mathbf{y}^s}(t)$  belong to  $\mathcal{H}_{t,+}^{\mathbf{y}}$  and hence,  $\mathcal{H}_t^{\mathbf{y}^s} \subseteq \mathcal{H}_{t,+}^{\mathbf{y}}$ . Since  $\mathbf{e}^s(t) = \mathbf{y}^s(t) - E[\mathbf{y}^s(t) \mid \mathcal{H}_t^{\mathbf{y}^s}]$ , this then implies that the components of  $\mathbf{e}^s(t)$  belong to  $\mathcal{H}_{t,+}^{\mathbf{y}}$ . Since  $\mathbf{z}_v^{\mathbf{y}}(t) = \sum_{i=1}^{n_\mu} \alpha_i \sqrt{p_i} \mathbf{z}_{vi}^{\mathbf{y}}(t+1)$ ,  $\mathbf{v}(t) = \sum_{i=1}^{n_\mu} \alpha_i \sqrt{p_i} \mathbf{z}_i^{\mathbf{y}}(t+1)$ , as  $\sum_{i=1}^{n_\mu} \alpha_i \mu_i = 1$ , it follows that  $\mathcal{H}_{t,+}^{\mathbf{y}} \subseteq \mathcal{H}_{t+1}^{\mathbf{y}}$  and from [11, Lemma 11] it follows that the components of  $\mathbf{z}_v^{\mathbf{y}^s}(t)$  belong to  $\mathcal{H}_t^{\mathbf{y}} \subseteq \mathcal{H}_{t,+}^{\mathbf{y}}$ . ■

**Lemma 5:** The entries of  $\{\mathbf{z}_v^{\mathbf{y}^s}(t), \mathbf{z}_v^{\mathbf{e}^s}(t)\}_{v \in \Sigma^+}, \mathbf{y}^s(t)$  and  $\mathbf{e}^s(t)$  are orthogonal to  $\mathcal{H}_{t,+}^{\mathbf{u}}$ .

*Proof:* The entries  $\mathbf{y}^s(t)$ ,  $\mathbf{e}^s(t)$ ,  $\{\mathbf{z}_v^{\mathbf{y}^s}(t), \mathbf{z}_v^{\mathbf{e}^s}(t)\}_{v \in \Sigma^+}$  belong to  $\mathcal{H}_{t,+}^{\mathbf{y}}$ , and by Lemma 1 the elements of  $\mathcal{H}_{t,+}^{\mathbf{y}}$  are orthogonal to  $\mathcal{H}_{t,+}^{\mathbf{u}}$ . ■

**Lemma 6:**  $\mathbf{r} = [(\mathbf{e}^s)^T, \mathbf{u}^T]^T$  is a white noise process w.r.t.  $\mu$  and  $E[\mathbf{e}^s(t) \mathbf{u}^T(t) \mu_\sigma^2(t)] = 0$  for all  $\sigma \in \Sigma$ .

*Proof:* We first show that  $\mathbf{r}$  is a ZMWSSI, by showing that  $\mathbf{r}$  satisfies the conditions of Definition 2 one by one.

**Part 1.** By assumption  $\mathbf{u}$  is a ZMWSSI and white noise process w.r.t.  $\mu$ . From the fact that  $\hat{\mathbf{S}}_s$  is an asGLSS of  $(\mathbf{y}^s, \mu)$  it follows that  $\mathbf{e}^s$  is also a ZMWSSI. Thus  $\mathbf{e}^s(t), \mathbf{u}(t), \{\mathbf{z}_w^{\mathbf{e}^s}(t), \mathbf{z}_w^{\mathbf{u}}(t)\}_{w \in \Sigma^+}$  are zero mean, square integrable, and hence so are  $\mathbf{r}(t)$  and  $\mathbf{z}_w^{\mathbf{r}}(t)$ . Finally, we show that  $\mathbf{r}(t), \mathbf{z}_w^{\mathbf{r}}(t), w \in \Sigma^+$  are jointly wide-sense stationary, i.e., for all  $s, t \in \mathbb{Z}$ ,  $E[\mathbf{h}_1(t)(\mathbf{h}_1(s)^T)]$ , where  $\mathbf{h}_1(\tau), \mathbf{h}_2(\tau) \in \{\mathbf{r}(\tau)\} \cup \{\mathbf{z}_w^{\mathbf{r}}(\tau)\}_{w \in \Sigma^+}$ , depends on  $t - s$ . We show only the case  $E[\mathbf{r}(t)(\mathbf{z}_w^{\mathbf{r}}(s))^T] = E[\mathbf{r}(t-s)(\mathbf{z}_w^{\mathbf{r}}(0))^T]$ ,  $t > s$ , the proof of the general case is similar. By repeated application of  $\mathbf{z}_v^{\mathbf{u}}(s) = \sum_{i=1}^{n_\mu} \alpha_i \mathbf{z}_{vi}^{\mathbf{u}}(s+1)$  it follows that the entries of  $\mathbf{z}_v^{\mathbf{u}}(s)$  belong to  $\mathcal{H}_{t,+}^{\mathbf{u}}$ . By Lemma 5,  $E[\mathbf{z}_w^{\mathbf{e}^s}(t)(\mathbf{z}_v^{\mathbf{u}}(s))^T] = 0$ , and hence  $E[\mathbf{z}_w^{\mathbf{r}}(t)(\mathbf{z}_v^{\mathbf{r}}(s))^T]$  is a block-diagonal matrix, and the diagonal blocks are  $E[\mathbf{z}_w^{\mathbf{e}^s}(t)(\mathbf{z}_v^{\mathbf{e}^s}(s))^T]$  and  $E[\mathbf{z}_w^{\mathbf{u}}(t)(\mathbf{z}_v^{\mathbf{u}}(s))^T]$ . As  $\mathbf{u}$  and  $\mathbf{e}^s$  are ZMWSSI, the latter blocks depend on  $t - s$ .

**Part 2.** By Lemma 4, the entries  $\mathbf{e}^s(t)$  belong to  $\mathcal{H}_{t,+}^{\mathbf{y}}$ . Moreover, by definition of sGLSS,  $\mathbf{w} = [\mathbf{v}^T, \mathbf{u}^T]^T$  is ZMWSSI. Hence, the  $\sigma$ -algebras  $\mathcal{F}_t^{\mathbf{w}}$  and  $\mathcal{F}_t^{\mathbf{u},+}$  are conditionally independent w.r.t.  $\mathcal{F}_t^{\mathbf{u},-}$ . Since  $\mathbf{e}^s(t)$  belongs to  $\mathcal{H}_{t,+}^{\mathbf{y}}$ ,  $\mathbf{e}^s(t)$  is measurable with respect to the  $\sigma$ -algebra generated by  $\{\mathbf{v}(t)\} \cup \{\mathbf{z}_v^{\mathbf{v}}(t)\}_{v \in \Sigma^+}$  and the latter  $\sigma$ -algebra is a subset of  $\mathcal{F}_t^{\mathbf{w}} \vee \mathcal{F}_t^{\mathbf{u},-}$ , where  $\mathcal{F}_1 \vee \mathcal{F}_2$  denotes the smallest  $\sigma$ -algebra generated by the union of sets  $\mathcal{F}_1$  and  $\mathcal{F}_2$ . That is,  $\mathbf{e}^s(t)$  is measurable w.r.t. the  $\sigma$  algebra  $\mathcal{F}_t^{\mathbf{w}} \vee \mathcal{F}_t^{\mathbf{u},-}$ . Hence,  $\mathcal{F}_t^{\mathbf{r}} \subseteq \mathcal{F}_t^{\mathbf{w}} \vee \mathcal{F}_t^{\mathbf{u},-}$ . Since  $\mathcal{F}_t^{\mathbf{w}}$  and  $\mathcal{F}_t^{\mathbf{u},+}$  are conditionally independent w.r.t.  $\mathcal{F}_t^{\mathbf{u},-}$ , from [18, Proposition 2.4] it follows that  $\mathcal{F}_t^{\mathbf{w}} \vee \mathcal{F}_t^{\mathbf{u},-}$  and  $\mathcal{F}_t^{\mathbf{u},+}$  are conditionally independent w.r.t.  $\mathcal{F}_t^{\mathbf{u},-}$ , and as  $\mathcal{F}_t^{\mathbf{r}} \subseteq \mathcal{F}_t^{\mathbf{w}} \vee \mathcal{F}_t^{\mathbf{u},-}$ , it follows that  $\mathcal{F}_t^{\mathbf{r}}$  and  $\mathcal{F}_t^{\mathbf{u},+}$  are conditionally independent w.r.t.  $\mathcal{F}_t^{\mathbf{u},-}$ .

Next we show that  $\mathbf{r}$  is a white noise process w.r.t.  $\mu$ . To this end, by [11, Lemma 7], it is enough to show that  $E[\mathbf{r}(t)(\mathbf{z}_w^{\mathbf{r}}(t))^T] = 0$ ,  $w \in \Sigma^+$ . By Lemma 5  $E[\mathbf{r}(t)(\mathbf{z}_w^{\mathbf{r}}(t))^T]$  is block diagonal, with the block on the diagonal being  $E[\mathbf{e}^s(t)(\mathbf{z}_w^{\mathbf{e}^s}(t))^T]$ ,  $E[\mathbf{u}(t)(\mathbf{z}_w^{\mathbf{u}}(t))^T]$ , and the latter are zero as  $\mathbf{e}^s$  and  $\mathbf{u}$  are white noise w.r.t.  $\mu$ .

Finally,  $E[\mathbf{e}^s(t) \mathbf{u}^T(t) \mu_\sigma^2(t)] = 0$ ,  $\sigma \in \Sigma$  follows from Lemma 5. ■

*Proof:* [Proof of Theorem 2] The proof is an extension

of [10, proof of Lemma 2]. From Lemma 6 it follows that  $\mathbf{e}^s$  and  $\mathbf{u}$  satisfy the condition of  $E[\mathbf{e}^s(t)(\mathbf{u}(t))^T \mu_\sigma^2(t)] = 0$ ,  $\sigma \in \Sigma$ . Since  $\hat{\mathbf{S}}_s$  and  $\hat{\mathbf{S}}_d$  are both asGLSS, it follows that the eigenevalues of  $\sum_{i=1}^{n_\mu} p_i \hat{A}_i \otimes \hat{A}_i$  are inside the open unit disk. Hence  $\hat{\mathbf{S}}$  satisfies the conditions of a sGLSS. Let  $\hat{\mathbf{x}}^s$  and  $\hat{\mathbf{x}}^d$  be the unique state processes of  $\hat{\mathbf{S}}_s$  and  $\hat{\mathbf{S}}_d$  respectively. Then  $\hat{\mathbf{x}}(t) = \hat{\mathbf{x}}^d(t) + \hat{\mathbf{x}}^s(t)$  is the unique state process of  $\hat{\mathbf{S}}$ . Indeed,  $\hat{\mathbf{x}}(t+1) = \sum_{i=1}^{n_\mu} (\hat{A}_i \hat{\mathbf{x}}(t) + \hat{B}_i \mathbf{u}(t) + \hat{K}_i \mathbf{e}^s(t)) \mu_i(t)$  holds and  $\hat{\mathbf{x}}(t)$  is a ZMWSSI, as it is a sum of two ZMWSSI processes. Finally, from  $\mathbf{y}^d(t) = \hat{C} \hat{\mathbf{x}}^d(t) + \hat{D} \mathbf{u}(t)$  (as  $\hat{\mathbf{S}}_d$  is an asGLSS of  $(\mathbf{y}^d, \mu)$ ) and  $\mathbf{y}^s(t) = \hat{C} \hat{\mathbf{x}}^s(t) + \mathbf{e}^s(t)$  (as  $\hat{\mathbf{S}}_s$  is an asGLSS of  $(\mathbf{y}^s, \mu)$ ), it follows that  $\mathbf{y}(t) = \hat{C} \hat{\mathbf{x}}(t) + \hat{D} \mathbf{u}(t) + \mathbf{e}^s(t)$ , i.e.,  $\hat{\mathbf{S}}$  is a sGLSS of  $(\mathbf{y}, \mathbf{u}, \mu)$ . ■

## REFERENCES

- [1] M.P. Balenzuela, A.G. Wills, C. Renton, and B. Ninness. Parameter estimation for jump markov linear systems. *Automatica*, 135:109949, 2022.
- [2] P. Billingsley. *Probability and measure*. Wiley, 1986.
- [3] O.L.V. Costa, M.D. Fragoso, and R.P. Marques. *Discrete-Time Markov Jump Linear Systems*. Springer Verlag, 2005.
- [4] P. Cox, M. Petreczky, and R. Tóth. Towards efficient maximum likelihood estimation of LPV-SS models. *Automatica*, 97(9):392–403, 2018.
- [5] P. Cox and R. Tóth. Linear parameter-varying subspace identification: A unified framework. *Automatica*, 123:109296, 2021.
- [6] T. Katayama. *Subspace Methods for System Identification*. Springer-Verlag, 2005.
- [7] F. Lauer and G. Bloch. *Hybrid System Identification: Theory and Algorithms for Learning Switching Models*. Springer, 2019.
- [8] A. Lindquist and G. Picci. *Linear Stochastic Systems: A Geometric Approach to Modeling, Estimation and Identification*. Springer Berlin, 2015.
- [9] M. Mejari and M. Petreczky. Realization and identification algorithm for stochastic LPV state-space models with exogenous inputs. In *3rd IFAC Workshop on Linear Parameter Varying Systems*, 2019.
- [10] M. Mejari and M. Petreczky. Realization and identification algorithm for stochastic LPV state-space models with exogenous inputs. *arXiv 1905.10113*, 2019.
- [11] M. Petreczky and R. Vidal. Realization theory for a class of stochastic bilinear systems. *IEEE Transactions on Automatic Control*, 63(1):69–84, 2018.
- [12] Giorgio Picci and Tohru Katayama. Stochastic realization with exogenous inputs and ‘subspace-methods’ identification. *Signal Processing*, 52(2):145–160, 1996.
- [13] D. Piga, P. Cox, R. Tóth, and V. Laurain. LPV system identification under noise corrupted scheduling and output signal observations. *Automatica*, 53:329–338, 2015.
- [14] E. Roupael, M. Petreczky, and L. Belkoura. On minimal LPV state-space representations in innovation form: an algebraic characterization. In *61st IEEE Conference on Decision and Control*, 2022.
- [15] Zh. Sun and Sh. S. Ge. *Switched linear systems : control and design*. Springer, London, 2005.
- [16] H. Tanaka and K. Ikeda. State estimation for closed-loop lpv system identification via kernel methods. *IFAC-PapersOnLine*, 56(2):11669–11674, 2023.
- [17] R. Tóth. *Modeling and Identification of Linear Parameter-Varying Systems*. Springer, 2010.
- [18] C. van Putten and J. H. van Schuppen. Invariance properties of the conditional independence relation. *Ann. Probab.*, 13(3):934–945, 1985.
- [19] J. W. van Wingerden and M. Verhaegen. Subspace identification of bilinear and LPV systems for open- and closed-loop data. *Automatica*, 45(2):372–381, 2009.
- [20] V. Verdult and M. Verhaegen. Subspace identification of multivariable linear parameter-varying systems. *Automatica*, 38(5):805–814, 2002.
- [21] V. Verdult and M. Verhaegen. Subspace identification of piecewise linear systems. In *43rd IEEE Conference on Decision and Control*, 2004.